



Andean alpaca population seems to be based on primary populations, also called primitive breeds, in which natural selection is favored over artificial selection. The mating system is panmictic and the phenotypic variation inside the flocks is high (Mason 1973; Denis 1982). For the moment, secondary standardized breeds do not seem to exist, apart from some experiences on very large private farms (Pacomarca, Mallkini). Nevertheless, the Andean alpaca population is subject to substantial genetic erosion through the diffusion of the full white phenotype. Suri is considered a phenotype with reduced fitness; for this reason, it is generally reared separately from Huacaya (Renieri *et al.*, 2009a).

### Population genetics by means of microsatellite analysis

A 10 microsatellite markers panel has been used by the ARI (Alpaca Registry Inc.) since 1998, mainly for parentage verification, and has found other applications throughout the years in other fields such as genome mapping, population structure and comparable genome analysis, as reviewed by Munyard *et al.* (2009). The continuous and ongoing effort in sequencing the alpaca genome will rapidly lead to a much larger set of markers, as demonstrated by a number of studies published in the past decade (Obreque *et al.*, 1999; Penedo *et al.*, 1999; McPartlan *et al.*, 1998). Reed and Chaves (2008) report an additional 1516 potential *loci* by blasting *bos taurus* SSRs and Munyard *et al.* (2009) have recently found a set of 9 tetranucleotide markers. Some of these markers have already been used to calculate genetic distances among different species of South American Camelids (Wheeler *et al.*, 2006; Bustamante *et al.*, 2002). Since Goldstein *et al.* (1995) evaluated their use for the calculation of genetic distances, they have been used in an increasingly large number of species.

Given the unusual post domestication evolutionary history of the species and the increasing interest in Suri fibre, the aim of this study was, therefore, to use a microsatellite panel to study the genetic distance between Suri and Huacaya alpacas and to assess the amplitude of genetic variability in the Peruvian alpaca population. The studied population belongs to a germplasm established approximately 20 years ago in Quimsachata (Puno province, Peru) and has been bred separately according to their coat type since then.

## Materials and methods

### Material collection and sample structure

Alpaca blood samples were collected in spring 2008 within a larger sampling program including skin biopsies and fibre. Animals were kept and managed in the Illpa-Puno Experimental Station in Quimsachata (Puno province, Peru) at an altitude of approximately 4200 m a.s.l. The animals belonged to two geographically separate and phenotypically pure flocks (one Suri and one Huacaya) that have been managed and bred separately since the alpaca germplasm was created at the experimental station 20 years ago. It is important to emphasize that common breeding practices in local Peruvian communities involve rearing Suri and Huacaya animals separately.

Blood samples were taken from a subset of 65 non-related animals selected for the microsatellite analysis. The sample was structured in order to include an equal number of Suri and Huacaya individuals and to respect the sex ratio as much as possible. It included

all the available males (n=15) and a subset of females (n=50): in total 32 Huacaya (7 males and 25 females) and 33 Suri (8 males and 25 females).

For convenience and due to the lack of basic facilities, blood was collected by spotting a total of 100 µL of blood on Whatman FTA Nucleic Acid Collection cards (# WB120205). Possibly due to the non-sterile and difficult conditions of facilities on the Peruvian plateau, not all samples allowed a sufficient quality/quantity of DNA to be amplified for all microsatellites, and in order to minimize the number of missing data, the dataset was rearranged to obtain a final number of 49 individuals: 10 males (5 Huacaya and 5 Suri) and 39 females (19 Huacaya and 20 Suri), and 13 microsatellites; LCA 19 was not included in the analysis.

### Microsatellite amplification details

All samples were processed in Italy and genomic amplification was carried out by LGS genetic laboratories (Cremona, Italy). The markers, dyes utilized, primer sequences, multiplex associations and allele sizes for the 14 microsatellites are shown in Table 1. The panel

**Table 1. Markers and multiplex reaction data.**

Marker	Fragment length	Alleles	Dye	5'-3' Primer pair sequences
LCA 19	80-122	17	Vic	taagtccagccccacactca ggtgaaggggctgatcttc
LCA 94	187-213	9	Pet	gtccatcatccagcacagg acatttggcaatctctggagaa
YWLL 44	84-136	18	Ned	ctcaacaatgctagacctgg gagaacacaggctggtgaata
YWLL 36	136-176	17	Vic	agtcttgggtggtgtagaa tgccaggatactgacagtgat
YWLL 43	128-164	10	Pet	Atacctctctctctctctc cctctacaacctgtagcca
YWLL 29	210-232	9	Fam	gaaggcaggagaaaaggtag cagaggcttaataactgcag
LCA 37	124-174	19	Fam	Aaacctaattacctcccca ccatgtagtgcaggacacg
LCA 5	178-218	13	Vic	Gtggttttgcccaagctc acctccagctcgggatttc
LCA 8	211-261	14	Pet	gctgaaccacaatgcaaa aatgcagatgctcctcagtt
LCA 65	159-193	14	Fam	Ttttcccctgtggtgaat aactcagctgtgtcagggg
LCA 66	216-266	24	Ned	gtgcagctccaatagtca ccagcatcgtccagattca
YWLL 40	176-190	7	Ned	cacatgacctgtccccttat ccagtgacagtgactaaga
LCA 99	263-297	11	Vic	caggtatcaggagacgggct agcatttatcaaggaaccaccg
YWLL 46	87-115	5	Fam	aagcagagtgatttaaccgtg ggatgactaagactgctctga

List of markers, allele fragment lengths, number of alleles, dyes used and respective primer pairs in the multiplex reactions.

has been optimized to be amplified in two multiplex reactions of 7 primer pairs each.

Amplifications were carried out in 20 µL reactions with final concentrations of 0.2 µM of each primer, 1X Ampli-Taq Gold Buffer, 1.5 mM MgCl<sub>2</sub>, 200 µM of each dNTP and a total of 2 units of Ampli-Taq Gold (#4398833, Applied Biosystems, Monza, Italy). After a first denaturation step at 94°C for 10 min, reactions were cycled 30 times (45 s at 94°C; 30 s at 58°C; 30 s at 72°C) with a final 10 min elongation step at 72°C.

### Software and statistical analysis

All microsatellite data were first checked with the software micro-checker (Oosterhout *et al.*, 2004) in order to spot null alleles and wrong size detections. The statistical analysis of the microsatellite data for the genetic variability measures, including the analysis of molecular variance (AMOVA) and the principal coordinate analysis (PCA), was performed using the latest version of the software Genalex 6.3 (Peakall and Smouse, 2006), while the Excel Microsatellite Toolkit (Park 2001) was used for calculating the polymorphism information content (PIC) for each allele.

A series of indices and parameters were calculated, such as allele frequencies and number of alleles, number of effective alleles, private alleles, expected and observed heterozygosity and fixation indices. The Fixation index (Wright's inbreeding coefficient) was calculated using the software Genalex as  $(H_e - H_o)/H_e$ . Arcos-Burgos and Muenke (2002) gives an exhaustive review of this measure which, when applied to microsatellite data, describes the probability that a given *locus* becomes fixed. Populations showing high levels of homozygosity will have a Fixation index significantly different from zero.

All these parameters were calculated for the whole dataset and for the two populations: Suri and Huacaya. In order to assess if there were statistically significant differences between these results, the null hypothesis was tested by means of one way analysis of variance (ANOVA) for observed heterozygosity (H<sub>o</sub>), expected heterozygosity (H<sub>e</sub>), Fixation index (F), number of effective alleles (N<sub>e</sub>) and PIC and by means of a non-parametric Mann-Whitney Test for number of alleles (N<sub>a</sub>).

The statistical population genetics package Fstat 2.9.3.2 (Goudet, 1995) was used to calculate deviation from Hardy-Weinberg equilibrium (HWE) and genotypic disequilibrium among *loci* applying a strict Bonferroni correction for multiple comparisons. The test for HWE and the test for linkage disequilibrium were carried out using 1300 randomizations:

1% table wide level was significant.

Cavalli-Sforza's chord distance and Reynolds-Weir Cockerham distance were calculated using Gendist, an application of the software package Phylip version 3.69 (Felsenstein, 1989). The first measure is assuming a stepwise mutation model in an infinite allele model with equilibrium between mutation and genetic drift, whereas the second and the third measures are dimensional models assuming only genetic drift.

Finally, the genetic structure of the sample was investigated with the Structure 2.3.3 software (Pritchard, 2000) and the identification of the most likely number of clusters (K) was made by the Evanno method (Evanno, 2005) using the online version of *Structure Harvester* (Earl, 2011). The burning period was set to 50,000 and repetitions of the MCMC chain to 106; the ancestry model chosen was the admixture model. Four replicates for each tested value of K (1-6) were performed.

## Results

### Hardy-Weinberg equilibrium and linkage disequilibrium

A number of indices have been calculated only on 12 microsatellites, excluding YWLL43, which is linked to the X sexual chromosome. Only one *locus* (LCA37) was found not to be in Hardy-Weinberg equilibrium after strict Bonferroni correction, showing excess of homozygosity ( $P < 0.05$ ). There was no differ-

ence in results when the sample was split into two different populations.

The test for genotypic disequilibrium between pairs of *loci* showed 9 *loci* associations out of 65 to be in some degree of linkage disequilibrium. The *loci* LCA8, LCA66 and LCA65 appear in 8 of the 9 associations showing linkage disequilibrium.

### Heterozygosity, polymorphism information content and fixation index

When the dataset was considered as a single population, the average H<sub>o</sub> for the 12 markers was high (H<sub>o</sub>=0.766; SE=0.044), extremely close to the average H<sub>e</sub> (H<sub>e</sub>=0.769; SE=0.033), and the unbiased expected heterozygosity (UHe=0.778; SE=0.033), with an overall average fixation index of 0.004 (SE 0.036). The high mean number of alleles (N<sub>a</sub>=9.667; SE=0.77), effective alleles (N<sub>e</sub>=4.89; SE=0.39) and the low fixation indices (F) confirm such high values of heterozygosity and genetic variability. The polymorphism information content (PIC) for each *locus* is in line with previous findings and ranges from 0.411 for *locus* YWLL46 to 0.826 for *locus* YWLL44. Table 2 shows H<sub>o</sub>, H<sub>e</sub>, UHe, PIC and F for each *locus* and as a mean for all *loci* with the relative standard errors.

There was no significant difference in results when the sample was analyzed as two separate populations (Huacaya and Suri). The null hypothesis was tested for H<sub>o</sub> (P=0.69), H<sub>e</sub> (P=0.61), F (P=0.95), N<sub>e</sub> (P=0.69), PIC (P=0.61) and N<sub>a</sub> (P=0.37). Genetic variability

Table 2. Population genetic parameters for a single population.

<i>Locus</i>	N	N <sub>a</sub>	N <sub>e</sub>	H <sub>o</sub>	H <sub>e</sub>	UHe	F	PIC
YWLL46	45	6 (5)	1.75	0.44	0.43	0.43	-0.03	0.41
LCA65	44	13 (14)	5.50	0.90	0.81	0.82	-0.11	0.79
YWLL40	46	6 (7)	4.80	0.82	0.79	0.80	-0.04	0.75
LCA5	49	7 (13)	3.90	0.83	0.74	0.75	-0.12	0.70
LCA66	48	13 (24)	6.10	0.89	0.83	0.84	-0.07	0.81
LCA8	49	9 (14)	6.18	0.83	0.83	0.84	0.01	0.81
LCA99	48	11 (11)	4.03	0.64	0.75	0.76	0.14	0.72
YWLL44	49	11(18)	6.41	0.77	0.84	0.85	0.08	0.82
LCA37	47	13 (19)	4.49	0.55	0.77	0.78	0.28	0.76
LCA94	47	7 (9)	4.09	0.68	0.75	0.76	0.10	0.72
YWLL36	49	10 (17)	6.16	0.91	0.83	0.84	-0.09	0.81
YWLL29	48	10 (9)	5.21	0.87	0.80	0.81	-0.08	0.78
YWLL43°	46	6 (10)	2.60	-	-	-	-	0.56
Mean	47.41	9.667	4.89	0.76	0.76	0.77	0.01	0.74
SE	0.48	0.77	0.39	0.04	0.03	0.03	0.03	0.03

N, number of individuals; N<sub>a</sub>, number of different alleles, in brackets values from previous studies (Obreque *et al.*, 1999; Penedo *et al.*, 1999; McPartlan *et al.*, 1998); N<sub>e</sub>, number of effective alleles; H<sub>o</sub>, observed heterozygosity; H<sub>e</sub>, expected heterozygosity; UHe, unbiased expected heterozygosity =  $[2N / (2N - 1)] * H_e$ ; F, fixation index; PIC, polymorphism information content; °Ho, H<sub>e</sub>, UHe and F not shown for X-linked YWLL43 *locus*.



parameters for the two populations are listed in Table 3. A number of private alleles were detected between the two phenotypes and a summary is given in Table 4 according to *loci* and phenotypes.

### Genetic distance, AMOVA, PCA and genetic structure

Genetic distance calculated by Pairwise Population Matrix of Nei's Genetic Distance and Unbiased Nei's Genetic Distance (Nei, 1978) were 0.062 and <0.0001, respectively. When calculated as the Cavalli-Sforza's chord distance (Cavalli-Sforza and Edwards, 1967) results showed 0.03. Reynolds-Weir Cockerham distance (Reynolds *et al.*, 1983), suggested to be more precise in the calculation of genetic distances between closely related species and breeds (Laval *et al.*, 2002), was 0.04.

The PCA calculated on the distance matrix among the individuals in Figure 1 graphically shows how samples from the Suri and Huacaya datasets overlap and do not segregate into different groups. The first 3 dimensions of the PCA explain 64.64% of the total variance. All the variance observed in the two populations with the AMOVA test, calculated both by *Fst* and *Rst* values, was due to variation within populations (100%) and not to variation between populations (0%). The software Structure rendered a maximum likelihood for *K*=2 both with the classic method by Pritchard (2000) and with that by Evanno (2005), therefore suggesting a possible dual ancestry for the analyzed sample. However, individuals from both phenotypes were assigned in equal proportions to the two clusters without a clear-cut distinction between the two groups, as shown in Figure 2.

## Discussion

In terms of genetic variability within the Peruvian alpaca sample analyzed, all parameters and findings, such as number of alleles (*Na*=9.667; *SE*=0.772), number of effective alleles (*Ne*=4.89; *SE*=0.388), observed and expected heterozygosity (*Ho*=0.766, *SE*=0.044; *He*=0.769, *SE*= 0.033), show that the Peruvian population is still conserving high genetic variability and does not show any sign of artificial selection pressure for the studied *loci*. The low fixation indices for these *loci* confirm this interpretation of the data (*F*=0.004; *SE*=0.036) and suggest that the microsatellite panel used is suitable for genetic diversity studies. In

**Table 3. Population genetic parameters for the two populations.**

Population	Locus	N	Na	Ne	Ho	He	UHe	F
Huacaya	YWLL46	21	5	1.42	0.33	0.29	0.30	-0.12
	LCA65	21	9	4.34	0.95	0.77	0.78	-0.23
	YWLL40	21	6	4.47	0.85	0.77	0.79	-0.10
	LCA5	24	6	4.08	0.91	0.75	0.77	-0.21
	LCA66	23	11	5.68	0.91	0.82	0.84	-0.10
	LCA8	24	8	6.36	0.83	0.84	0.86	0.01
	LCA99	23	10	4.14	0.65	0.75	0.77	0.14
	YWLL44	24	9	5.78	0.70	0.82	0.84	0.14
	LCA37	24	9	4.15	0.50	0.76	0.77	0.34
	LCA94	24	6	3.61	0.62	0.72	0.73	0.13
	YWLL36	24	9	6.29	0.91	0.84	0.85	-0.09
	YWLL29	23	9	5.23	0.82	0.80	0.82	-0.02
	YWLL43§	22	5	2.28	-	-	-	-
	Suri	YWLL46	24	5	2.11	0.54	0.52	0.53
LCA65		23	11	5.71	0.87	0.82	0.84	-0.05
YWLL40		25	6	4.92	0.80	0.79	0.81	-0.01
LCA5		25	6	3.37	0.76	0.70	0.71	-0.08
LCA66		25	9	6.28	0.88	0.84	0.85	-0.04
LCA8		25	9	5.68	0.84	0.82	0.84	-0.01
LCA99		25	8	3.85	0.64	0.74	0.75	0.13
YWLL44		25	10	6.06	0.84	0.83	0.85	-0.01
LCA37		23	13	4.76	0.60	0.79	0.80	0.23
LCA94		23	7	4.40	0.73	0.77	0.79	0.04
YWLL36		25	8	5.98	0.92	0.83	0.85	-0.10
YWLL29		25	10	5.04	0.92	0.80	0.81	-0.14
YWLL43°		24	5	2.75	-	-	-	-
		N	Na	Ne	Ho	He	UHe	F
Huacaya	Mean	23	8.08	4.63	0.75	0.74	0.76	-0.01
	SE	0.36	0.54	0.39	0.05	0.04	0.04	0.05
Suri	Mean	24.41	8.50	4.85	0.78	0.77	0.79	-0.01
	SE	0.26	0.66	0.36	0.03	0.02	0.02	0.03

N, number of individuals; Na, number of different alleles; Ne, number of effective alleles; Ho, observed heterozygosity; He, expected heterozygosity; UHe, unbiased expected heterozygosity; F, fixation index; °Ho, He, UHe and F not shown for X-linked YWLL43 locus.

comparison with the data collected and analyzed by the ARI since 1998 and by several other authors, the sample in this study shows the presence of a high number of alleles, matching the whole allelic range described in previous publications (Obreque *et al.*, 1999; Penedo *et al.*, 1999; McPartlan *et al.*, 1998). Values for the polymorphic information content (*PIC*=0.746, *SE*=0.033) were also in line with previous findings. Only one *locus* (YWLL46) showed a *PIC* value less than 0.7 (*PIC*=0.411), which reflects the lower than average *Na* and *Ne* found for this specific *locus* (*Na*=6; *Ne*=1.753) in previous studies. Nevertheless, this value is higher than that from previous bibliographic data (Lang *et al.*, 1996).

There was no significant difference between the two phenotypic groups when these parameters were evaluated separately. In terms of genetic distance and differentiation between the two phenotypes the PCA

**Table 4. List of private alleles in each population.**

Population	Locus	Alleles	
Huacaya	YWLL46	95	
	LCA65	175, 179	
	LCA5	190	
	LCA66	220, 244, 252, 256	
	LCA99	278, 290, 294	
	YWLL44	108	
	YWLL43	148	
	YWLL36	168, 172	
	Suri	YWLL46	111
		LCA65	165, 177, 185, 189
LCA5		192	
LCA66		240, 260	
LCA8		245	
LCA99		268	
YWLL44		116, 134	
YWLL43		162	
LCA37		130, 138, 148, 170	
LCA94		197	
YWLL36	170		
YWLL29	238		

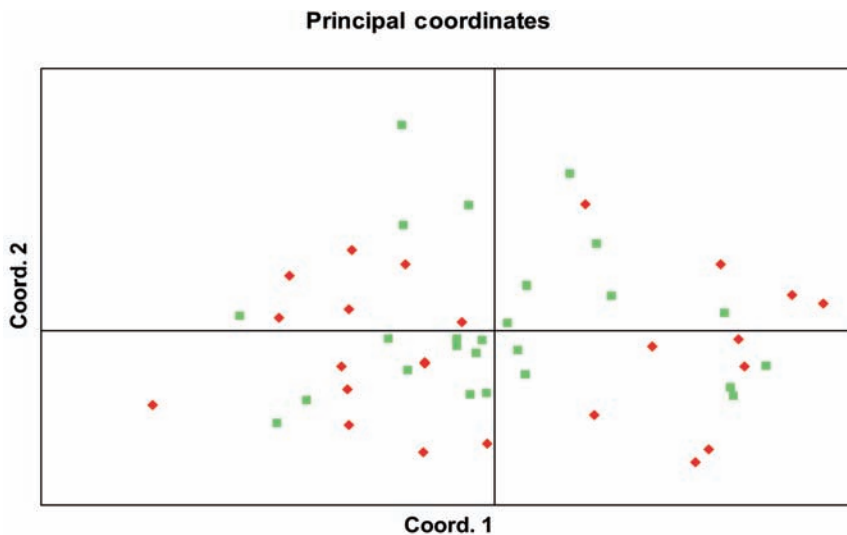


Figure 1. Principal coordinates analysis of Suri and Huacaya based on the distance matrix from molecular data. Red, Huacaya; green, Suri.

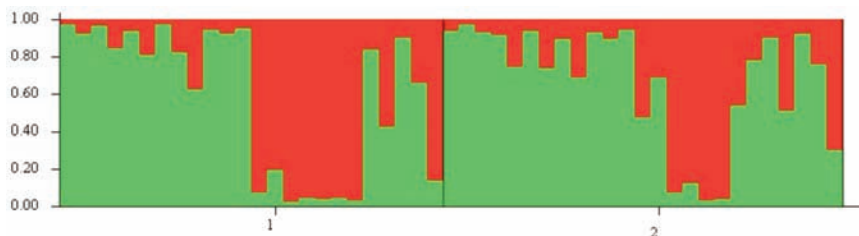


Figure 2. Bar plot for cluster assignment. 1, Huacaya; 2, Suri. Colored bars indicate the probability of assignment to either cluster 1 (green) or cluster 2 (red).

analysis shows no separate segregation or grouping of Suri and Huacaya individuals.

To our knowledge, no previous studies have investigated genetic distance between Suri and Huacaya alpaca by means of codominant markers, although microsatellite markers have been used in the past to investigate the phylogeny of a wider group of South American Camelids (Wheeler *et al.*, 2006, Bustamante *et al.*, 2002).

When ANOVA was carried out considering the two phenotypes as two separate populations, it clearly identified the source of all variance in the component *within populations*, excluding any source of variance to be found between populations. This result is supported by Nei's index of genetic distance, Cavalli-Sforza's chord distance and Reynolds-Weir Cockerham distance, which also show no differentiation between the two populations. There are two important factors to be taken into consideration while interpreting these results. Firstly, the germplasm established 20 years ago at the experimental station of Quimsachata is not subjected to genetic selection and was created

for the sole purpose of conserving the genetic diversity of the species. Secondly, although Suri and Huacaya alpacas at the experimental station have been bred and managed separately since the creation of the germplasm, the time interval of 20 years is unlikely to generate genetic differentiation between the two phenotypes, especially considering the absence of selection, the reproductive physiology of the species and its generation time (Mason, 1973).

## Conclusions

Given these considerations, if a secondary breed structure had been present within the species at the time of the creation of the germplasm, it would have been preserved by the breeding practices in place at the experimental station. Nevertheless, the data obtained from the 13 *loci* suggest no genetic divergence between the two phenotypes and do not support the idea of two distinct populations of Peruvian Suri and Huacaya alpacas.

Furthermore, the two phenotypes have similar genetic parameters in terms of allelic frequencies and genetic variability, showing high values both in terms of allelic richness and heterozygosity.

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